

Derivation of the chi-squared distribution

Let us have a spherically symmetrical probability density function

$$f_{\mathbf{x}}(x_1, x_2, \dots, x_n) = p(\sqrt{x_1^2 + x_2^2 + \dots + x_n^2}) \quad (1)$$

everywhere in \mathbb{R}^n , where $r = \sqrt{x_1^2 + x_2^2 + \dots + x_n^2}$ and $p(r)$ is a nonnegative function of the single variable r . To subject this joint density function to the normalization condition

$$\int_{-\infty}^{\infty} \int_{-\infty}^{\infty} \dots \int_{-\infty}^{\infty} f_{\mathbf{x}}(x_1, x_2, \dots, x_n) dx_1 dx_2 \dots dx_n = 1 \quad (2)$$

we must evaluate its integral over all of \mathbb{R}^n . We can do so by utilizing the fact that the probability density function $f_{\mathbf{x}}(x_1, x_2, \dots, x_n)$ is constant over a hyperspherical shell. The volume dV of a hyperspherical shell of radius r and thickness dr in n -dimensional space is proportional to $r^{n-1}dr$

$$dV = k_n r^{n-1} dr \quad (3)$$

Hence

$$\int_{-\infty}^{\infty} \int_{-\infty}^{\infty} \dots \int_{-\infty}^{\infty} f_{\mathbf{x}}(x_1, x_2, \dots, x_n) dx_1 dx_2 \dots dx_n = k_n \int_0^{\infty} r^{n-1} p(r) dr \quad (4)$$

$$= 1$$

The constant k_n can be worked out by picking a particular joint density function for which both sides of equation (4) can be easily evaluated. A good function for that evaluation is a Gaussian function

$$f_{\mathbf{x}}(x_1, x_2, \dots, x_n) = (2\pi)^{-n/2} e^{-(x_1^2 + x_2^2 + \dots + x_n^2)/2} \quad (5)$$

for which

$$\int_{-\infty}^{\infty} \int_{-\infty}^{\infty} \dots \int_{-\infty}^{\infty} e^{-(x_1^2 + x_2^2 + \dots + x_n^2)/2} dx_1 dx_2 \dots dx_n \quad (6)$$

$$= \int_{-\infty}^{\infty} e^{-x_1^2/2} dx_1 \int_{-\infty}^{\infty} e^{-x_2^2/2} dx_2 \dots \int_{-\infty}^{\infty} e^{-x_n^2/2} dx_n = (2\pi)^{n/2}$$

and after changing the variables of integration

$$k_n \int_0^{\infty} r^{n-1} e^{-r^2/2} dr = (2\pi)^{n/2} \quad (7)$$

and if

$$f_{\mathbf{x}}(x_1, x_2, \dots, x_n) = (2\pi)^{-n/2} e^{-(x_1^2 + x_2^2 + \dots + x_n^2)/2} \quad (8)$$

then

$$k_n (2\pi)^{-n/2} \int_0^\infty r^{n-1} e^{-r^2/2} dr = 1 \quad (9)$$

We now compute the integral in equation (9)

$$\begin{aligned} \int_0^\infty r^{n-1} e^{-r^2/2} dr &= \left| \begin{array}{l} r^2/2 = u \\ r dr = du \\ r = (2u)^{1/2} \end{array} \right| = \int_0^\infty r^{n-2} e^{-r^2/2} r dr \\ &= \int_0^\infty (2u)^{(n-2)/2} e^{-u} du = \frac{1}{2} 2^{n/2} \int_0^\infty u^{n/2-1} e^{-u} du = \frac{1}{2} 2^{n/2} \Gamma(n/2) \end{aligned} \quad (10)$$

and we have

$$k_n (2\pi)^{-n/2} \frac{1}{2} 2^{n/2} \Gamma(n/2) = 1 \quad (11)$$

what gives

$$k_n = \frac{2\pi^{n/2}}{\Gamma(n/2)} \quad (12)$$

and what we can rewrite as

$$f_{\mathbf{z}} = 2^{-(n/2)+1} [\Gamma(n/2)]^{-1} z^{n-1} e^{-z^2/2} \quad (13)$$

for

$$\mathbf{z} = (\mathbf{x}_1^2 + \mathbf{x}_2^2 + \dots + \mathbf{x}_n^2)^{1/2} \quad (14)$$

Now we introduce the variable

$$\mathbf{w} = \mathbf{z}^2 = \mathbf{x}_1^2 + \mathbf{x}_2^2 + \dots + \mathbf{x}_n^2 \quad (15)$$

and the question is what is the probability distribution function $f_{\mathbf{w}}(\mathbf{w})$ if we know the function $f_{\mathbf{z}}(\mathbf{z})$. We just need to substitute

$$\begin{aligned} w &= z^2 \\ z &= w^{1/2} \\ 2z dz &= dw \end{aligned} \quad (16)$$

in

$$f_{\mathbf{z}} dz = 2^{-(n/2)+1} [\Gamma(n/2)]^{-1} z^{n-1} e^{-z^2/2} dz \quad (17)$$

This gives

$$\begin{aligned} f_{\mathbf{z}} dz &= 2^{-n/2} [\Gamma(n/2)]^{-1} z^{n-2} e^{-z^2/2} 2z dz \\ &= 2^{-n/2} [\Gamma(n/2)]^{-1} w^{(n-2)/2} e^{-w/2} dw = f_{\mathbf{w}} dw \end{aligned} \quad (18)$$

The function

$$f_{\mathbf{w}}(w) = 2^{-n/2} [\Gamma(n/2)]^{-1} w^{(n-2)/2} e^{-w/2} \quad (19)$$

is the probability distribution function of the random variable \mathbf{w} which is called *chi-squared* (χ^2) and (19) is called the chi-squared distribution. We notice that the chi-squared distribution is defined for the random variable \mathbf{w} having values $w \geq 0$ defined in equation (15).

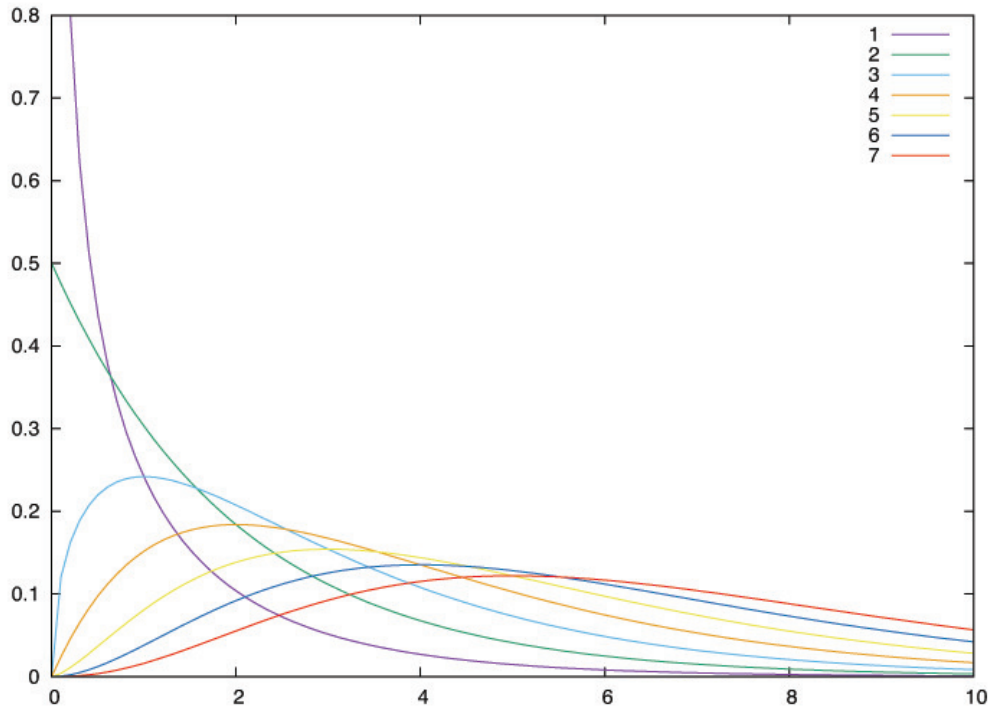


Figure 1: Plots of chi-squared probability density functions for $n = 1, 2, \dots, 7$

References

- [1] Helstrom, Carl W. (1984) *Probability and stochastic processes for engineers* Macmillan Publishing Company New York, Collier Macmillan Publishers London

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